

# Physics 403

Unfolding

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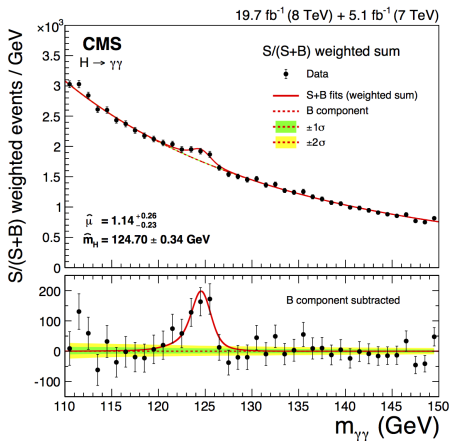
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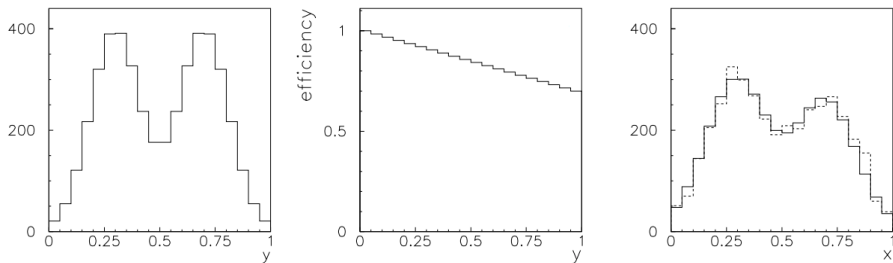
# Motivation



- One of the most common tasks in physics is the publication of spectra, by which we mean a **binned histogram** showing the distribution of events in some observable quantity
- Example: diphoton mass spectrum  $m_{\gamma\gamma}$  from CMS [1] showing the **Higgs resonance**
- **Problem:** spectra are often **smeared** and **distorted** by the finite resolution and thresholds of your apparatus
- Unfolding is used to fix these distortions

## Example: A Bin-Dependent Instrumental Response

Suppose we have a “true” spectrum given by the histogram on the left [2]



The center plot shows efficiency as a function of bin. E.g., the instrumental efficiency **decreases** as a function of the parameter  $y$

The right plot shows the measured spectrum (dashed) and expectation from Monte Carlo (solid) showing **distortion** in the counts

## Unfolding in 2D

- Note that this problem also applies to 2D data such as binned images



- We can try to **deconvolve** smearing and blur from an image [3]
- What's required is some model of the smearing effect. By “smearing” we refer to an effect that results in an event being classified or reconstructed in the **wrong bin**

# Forward Folding and Unfolding

Smearred data can be analyzed in a couple of ways:

1. **Forward Folding:** take a **theoretical spectrum**, smear it, and then compare the result to the data. The best fit gives you the **true spectrum**
2. **Unfolding:** take an observation which has smearing and other detector effects and try to **deconvolve** those effects

Forward folding is **considerably easier** than unfolding, so when possible it's best to do that. This is most appropriate when you just want to compare data from one experiment to a theoretical prediction

However, sometimes it is necessary to unfold your data. For example, if you want to **compare spectra** across several experiments, you need to remove the instrument response function to make unambiguous comparisons

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# Unfolding Formalism

Define the following terms (following the notation of [4]):

- ▶ **Truth:** a “truth” spectrum  $\mathbf{T} = (T_1, T_2, \dots, T_{N_t})$  represents the binned counts that would be observed without smearing
- ▶ Truth spectra are usually estimated with Monte Carlo. Let the true counts be  $\hat{\mathbf{T}}$  and the **Monte Carlo truth** be  $\tilde{\mathbf{T}}$ . Ideally,  $\tilde{\mathbf{T}} = \hat{\mathbf{T}}$
- ▶ **Reco:** a “reco” spectrum  $\mathbf{R} = (R_1, R_2, \dots, R_{N_r})$  is the number of events **expected to be reconstructed** in a bin
- ▶ **Data:** a “data” spectrum  $\mathbf{D} = (D_1, D_2, \dots, D_{N_d})$  is the number of events **observed** in a bin after smearing. We expect  $\mathbf{D}$  to follow a Poisson distribution with mean  $\mathbf{R}$
- ▶ **Migration matrix:** a matrix  $\mathcal{M}_{tr}$  defined by the joint PDF  $p(t, r)$  of an event being produced in true bin  $t$  and reconstructed in bin  $r$
- ▶ **Response matrix:** a matrix  $\mathcal{P}_{tr}$  defined by the conditional probability  $p(r|t)$



# The Unfolding Problem

- Fundamentally, the unfolding problem requires us to calculate

$$p(\mathbf{T}|\mathbf{D}, \mathcal{M}) = \frac{1}{Z} \mathcal{L}(\mathbf{D}|\mathbf{T}, \mathcal{M}) \pi(\mathbf{T}, \mathcal{M})$$

- If the data follow a Poisson distribution, then

$$\mathcal{L}(\mathbf{D}|\mathbf{T}, \mathcal{M}) = \prod_{r=1}^{N_r} \frac{R_r^{D_r}}{D_r!} e^{-R_r}$$

where the reconstructed counts are related to the **true counts** by

$$R_r = \sum_{t=1}^{N_t} T_t \cdot p(r|t)$$

- Recalling that  $p(r|t)$  is the probability of reconstructing an event in bin  $r$  given that it should have been in bin  $t$  before smearing, we have

$$p(r|t) = \frac{p(t, r)}{p(t)} = \frac{\mathcal{M}_{tr}}{\epsilon_t^{-1} \sum_{k=1}^{N_r} \mathcal{M}_{tk}}, \quad \epsilon_t = \frac{\sum_{r=1}^{N_r} \mathcal{M}_{tr}}{p(t)}$$

# Accounting for the Presence of Background

- ▶ It is typical for the data to be contaminated by sources of **background**
- ▶ Example: in gamma-ray experiments, the cosmic-ray background reduction efficiency is typically 99.9%, but since the cosmic-ray flux is  $1000\times$  larger than the gamma-ray flux, the resulting signal/noise ratio is 1:1
- ▶ We can account for background by adding it to the **expectation of the reconstructed counts**:

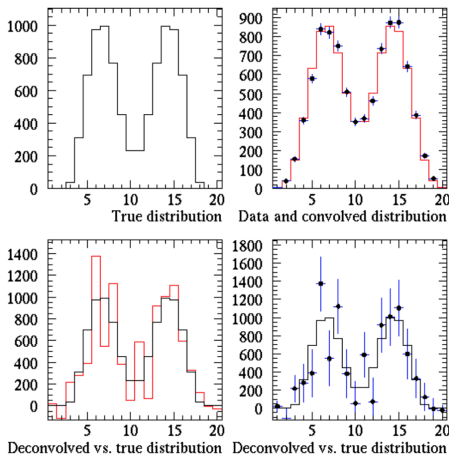
$$R_r = B_r + \sum_{t=1}^{N_t} T_t \cdot p(r|t)$$

where  $B_r$  is the number of background events in bin  $r$ . In matrix notation this is written

$$\mathbf{R} = \mathbf{B} + \mathcal{P}^T \mathbf{T}$$

# The Noise Amplification Problem

- ▶ Unfolding involves inverting the migration matrix to go from observed counts  $\mathbf{D}$  to true counts  $\mathbf{T}$
- ▶ Unfortunately this procedure tends to **amplify noise** in the data
- ▶ Right: 3-10% random scatter in the data is amplified to  $> 20\%$  errors in unfolded counts (from Oser)
- ▶ The off-diagonal terms in  $\mathcal{P}^T \mathbf{D}$  may be large, causing **oscillations** (zig-zagging) in the unfolding



# Regularization

- ▶ The maximum likelihood estimator (MLE) of  $\mathbf{T}$  is  $\mathcal{P}^\top \mathbf{D}$
- ▶ The MLE is the unbiased estimator with the **smallest possible variance**. Unfortunately the variance is still huge!
- ▶ Solution: introduce a bias into the estimator. The fit is now worse (w.r.t. the likelihood) but is smoothed by an amount that we specify
- ▶ Introduce a **regularization parameter**  $S(\mathbf{T})$  that increases as  $\mathbf{T}$  oscillates, and maximize

$$\tilde{\mathcal{L}}(\mathbf{T}) = \mathcal{L}(\mathbf{D}|\mathbf{T}) \cdot e^{-\alpha \cdot S(\mathbf{T})}$$

- ▶ Essentially we are defining the prior on  $\mathbf{T}$  as

$$\pi(\mathbf{T}) = e^{-\alpha \cdot S(\mathbf{T})}$$

If the prior is constant then  $p(\mathbf{T}|\mathbf{D})$  may be too wide, so different  $\mathbf{T}$ 's are equally likely. So we penalize certain solutions with  $S(\mathbf{T})$

# Tikhonov Regularization

- ▶ There is freedom to define  $\alpha$  and the regularization function, which determines the smoothness of the unfolded counts
- ▶ **Tikhonov regularization** is a common approach for defining smoothing functions. We define

$$S(f) = - \int dx \left( \frac{d^k f}{dx^k} \right)^2$$

where  $k$  represents the  $k^{\text{th}}$ -order derivative of  $f$  and  $f$  is the deconvolved distribution

- ▶ Common approach: set  $k = 2$ , which penalizes curvature in  $f$  (nonzero second derivatives)
- ▶ If we don't want to just favor linear functions of  $f$  we can use higher-order derivatives, or some linear combination of derivatives

# Tikhonov Regularization

- ▶ In a binned spectrum, if we want to penalize the **curvature** between bins, we would write

$$S(\mathcal{T}) = \sum_{t=2}^{N_t-1} (\Delta_{t+1,t} - \Delta_{t,t-1})^2$$

where

$$\Delta_{t_1,t_2} = T_{t_1} - T_{t_2}$$

- ▶ One can also try to **penalize variations in the first derivative** and account for varying bin sizes:

$$S(\mathcal{T}) = \sum_{t=2}^{N_t-1} \frac{|\delta_{t+1,t} - \delta_{t,t-1}|}{|\delta_{t+1,t} + \delta_{t,t-1}|}$$

where  $\delta_{t_1,t_2}$  is related to the bin width  $w_t$  and bin center  $c_t$  by

$$\delta_{t_1,t_2} = \frac{T_{t_1}/w_{t_1} - T_{t_2}/w_{t_2}}{c_{t_1} - c_{t_2}}, \quad w_t = m_t - m_{t-1}, \quad c_t = (m_t + m_{t-1})/2$$

# Maximum Entropy Regularization

- ▶ Without any prior knowledge about the distribution of  $\mathbf{T}$  in the bins, a reasonable choice for  $S$  is the **maximum entropy**

$$S(\mathbf{T}) = - \sum_{t=1}^{N_t} \frac{T_t}{\sum T_t} \ln \frac{T_t}{\sum T_t}$$

- ▶ Recall that the maximum entropy distribution is the one that favors the **largest possible value for the bin multiplicity**
- ▶ This means that it favors a relatively flat distribution, since entropy tends to be maximized when the bin counts are relatively equal
- ▶ If you don't want a flat distribution, you shouldn't use the maximum entropy. Instead, you can try **cross-entropy**

$$S(\mathbf{T}) = - \sum_{t=1}^{N_t} \frac{T_t}{\sum T_t} \ln \frac{T_t}{q_t \sum T_t}$$

where  $q$  contains any prior knowledge you may have about the true distribution. If  $q_i = 1/N_t$  this reduces back to the entropy

# Regularization using Monte Carlo

- ▶ If you really trust your Monte Carlo  $\tilde{\mathbf{T}}$  you could set your prior to [4]

$$\pi(\mathbf{T}) = \prod_{t=1}^{N_t} \exp \left[ -\frac{(\mathbf{T}_t - \tilde{\mathbf{T}}_t)^2}{2(\tilde{\mathbf{T}}_t/\alpha)^2} \right]$$

- ▶ Here the prior is proportional to a multivariate Gaussian with **no correlations** between bins
- ▶ The effect of the prior is to disfavor  $\mathbf{T}$  far from  $\tilde{\mathbf{T}}$
- ▶ The free parameter  $\alpha$  adjusts the width of the Gaussian. Larger values of  $\alpha$  imply a stricter constraint, forcing the unfolded counts to match the result of the Monte Carlo



## Choosing a Regularization Parameter

There is no recipe for choosing a regularization parameter, but you can pick one of several criteria [2, 5]:

1. Minimize the **mean squared error** (MSE):

$$MSE = \frac{1}{N_t} \sum_{t=1}^{N_t} \text{var}(T_t) + \hat{b}_t^2, \quad \hat{b}_t = E(T_t) - T_t$$

2. Tune  $\alpha$  and  $S$  so that for each bin

$$2\Delta \ln \mathcal{L} = 2(\ln \mathcal{L}_{\max} - \ln \mathcal{L}) = N_t$$

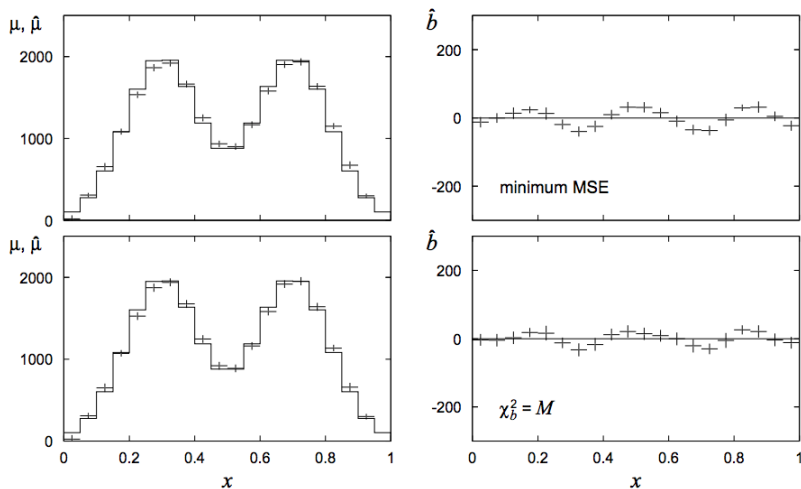
3. Tune  $\alpha$  and  $S$  so that the biases and variances are balanced:

$$\chi_b^2 = \sum_{t=1}^{N_t} \frac{\hat{b}_t^2}{\text{var}(\hat{b}_t)} \approx N_t$$

There is a trade-off between **bias** and **variance** that you have to choose for your application

# Example: Unfolding with Tikhonov Regularization

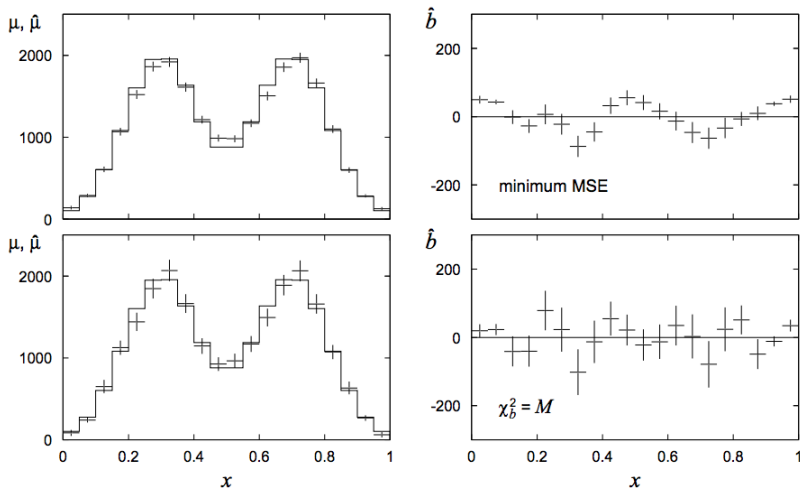
Unfolded distributions using Tikhonov regularization [2]



The parameter  $\alpha$  was tuned using the MSE and  $\chi_b^2 = N_t$

# Example: Unfolding with Maximum Entropy Regularization

Unfolded distributions using MaxEnt regularization [2]



The parameter  $\alpha$  was tuned using the MSE and  $\chi_b^2 = N_t$

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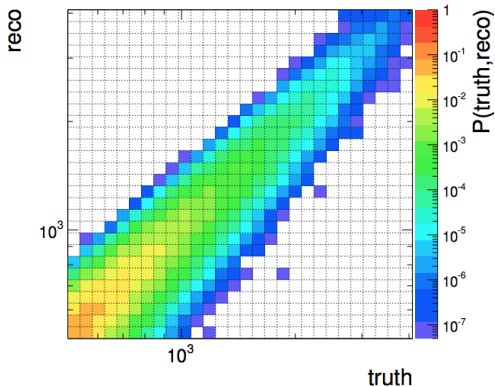
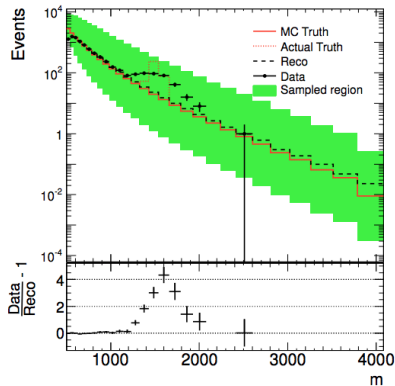
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## Example: Steeply Falling Spectrum with a Bump

A steeply falling spectrum with a bump, from [4]



The true spectrum has a bump. The Monte Carlo truth does not. The data are smeared by the bin-to-bin **migration matrix** shown at right

# MCMC Sampling

The procedure for unfolding is as follows:

- ▶ Choose a prior for  $\mathbf{T}$
- ▶ Sample the  $N_t$ -dimensional **hypercube** using MCMC
- ▶ For each bin, find the mode (maximum) of the posterior  $p(\mathbf{T}|\mathbf{D})$

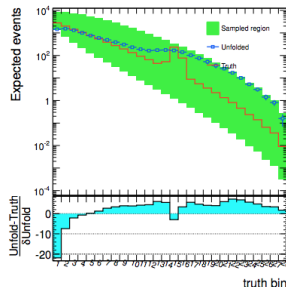
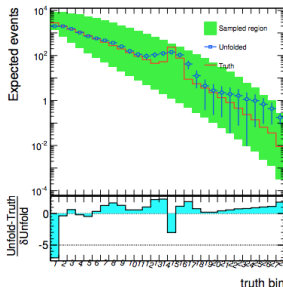
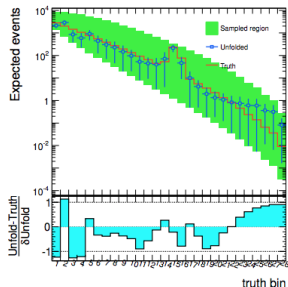
It's good practice to try different smoothing penalties  $S(\mathbf{T})$  and smoothing factors  $\alpha$

Usual caveats about MCMC: only use data after **burn-in**, and plot the marginal distributions of  $\mathbf{T}$  to see if they are unimodal

# Unfolding with a MaxEnt Penalty

The unfolded spectrum is reconstructed using **maximum entropy regularization**

$$S(\mathcal{T}) = - \sum_{t=1}^{N_t} \frac{T_t}{\sum T_t} \ln \frac{T_t}{\sum T_t}$$

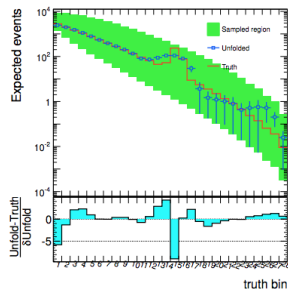
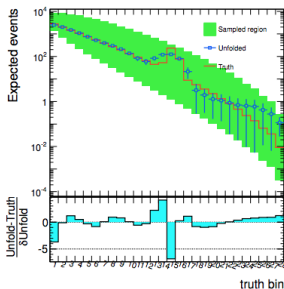
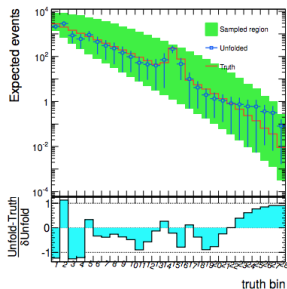


Left: no penalty ( $\alpha = 0$ ). Middle:  $\alpha = 10^3$ . Right:  $\alpha = 3 \times 10^3$ .

# Unfolding with a Curvature Penalty

The unfolded spectrum is reconstructed using the **curvature penalty**

$$S(\mathcal{T}) = \sum_{t=2}^{N_t-1} (\Delta_{t+1,t} - \Delta_{t,t-1})^2$$



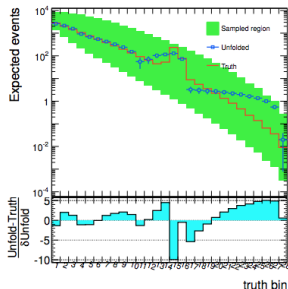
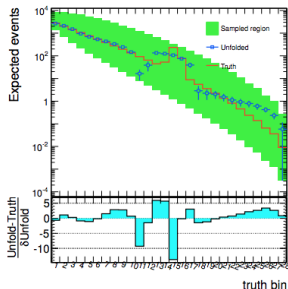
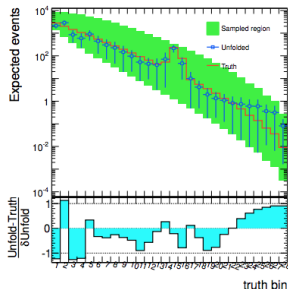
Left: no penalty ( $\alpha = 0$ ). Middle:  $\alpha = 3 \times 10^{-4}$ . Right:  $\alpha = 6 \times 10^{-4}$ .



# Unfolding Accounting for Nonuniform Bins

The unfolded spectrum is reconstructed using the **penalty**

$$S(\mathbf{T}) = \sum_{t=2}^{N_t-1} \frac{|\delta_{t+1,t} - \delta_{t,t-1}|}{|\delta_{t+1,t} + \delta_{t,t-1}|}$$

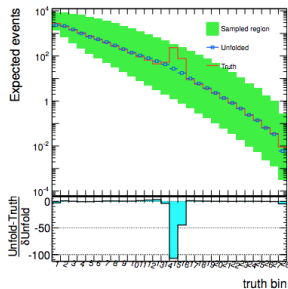
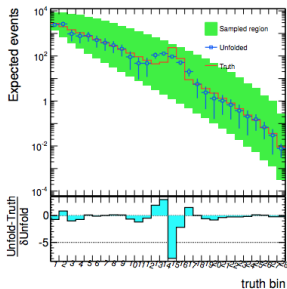
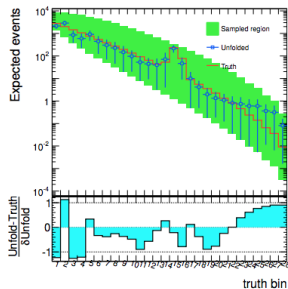


Left: no penalty ( $\alpha = 0$ ). Middle:  $\alpha = 10$ . Right:  $\alpha = 20$ .

# Unfolding with Gaussian Regularization

The unfolded spectrum is reconstructed using the **prior**

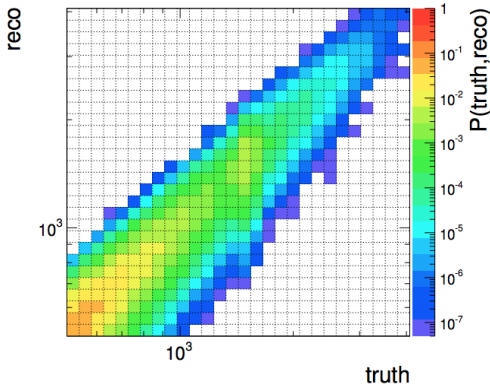
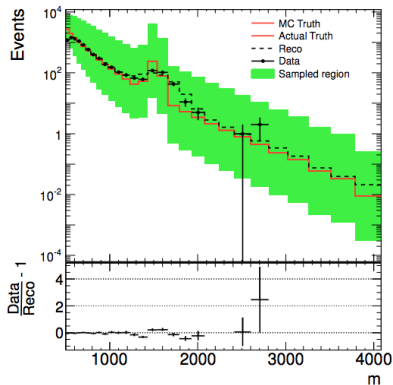
$$\pi(\mathbf{T}) = \prod_{t=1}^{N_t} \exp \left[ -\frac{(\mathcal{T}_t - \tilde{\mathcal{T}}_t)^2}{2(\tilde{\mathcal{T}}_t/\alpha)^2} \right]$$



Left: no penalty ( $\alpha = 0$ ). Middle:  $\alpha = 1$ . Right:  $\alpha = 10$ .

## Example: Steeply Falling Spectrum with an Expected Bump

A steeply falling spectrum with a bump, from [4]

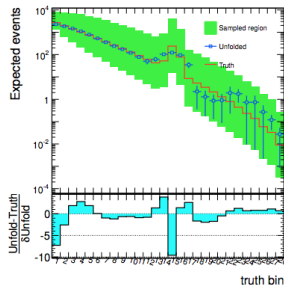
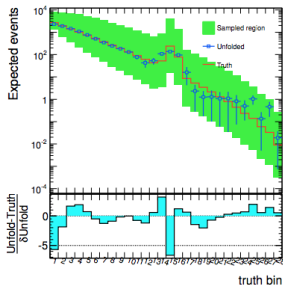
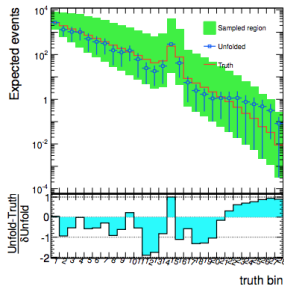


This time, the **bump is expected and included in the MC truth**. Is there any difference in the unfolding?

# Unfolding with a Curvature Penalty

The unfolded spectrum is reconstructed using the **curvature penalty**

$$S(\mathcal{T}) = \sum_{t=2}^{N_t-1} (\Delta_{t+1,t} - \Delta_{t,t-1})^2$$



No real improvement in appearance of the bump w.r.t. case where  $\tilde{\mathcal{T}}$  did not contain a bump

# Summary

- ▶ Unfolding is a technique used to remove instrumental **smearing** and **efficiency** artifacts from a binned spectrum
- ▶ After unfolding, the unbiased **maximum likelihood estimator** tends to have big variances which show up as zig-zagging between neighboring bins
- ▶ The fix for oscillations is to apply a **smoothing function** that penalizes zig-zagging. There is a lot of freedom in how to do this
- ▶ There is a trade off between the bias in the estimator and the variance. You have to decide what is appropriate; there is no recipe
- ▶ Best approach: **run a data challenge** to see if the kind of effect you are looking for is washed out by how you unfold
- ▶ Cowan suggests several **figures of merit** for balancing variance and bias [2, 5] that are good starting points for this kind of analysis

# References I

- [1] Vardan Khachatryan et al. “Observation of the diphoton decay of the Higgs boson and measurement of its properties”. In: *Eur.Phys.J.* C74.10 (2014), p. 3076. arXiv: 1407.0558 [hep-ex].
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